

## OPTIMAL IMPORTANCE SAMPLING FOR SIMULATION OF LÉVY PROCESSES

Guangxin Jiang

Michael C. Fu

Department of Economics and Finance  
City University of Hong Kong  
Kowloon, Hong Kong

Robert H. Smith School of Business and  
Institute for Systems Research, University of Maryland  
College Park, MD, 20742, USA

Chenglong Xu

Department of Mathematics  
Tongji University  
Shanghai, 200092, China

### ABSTRACT

This paper provides an efficient algorithm using Newton's method under sample average approximation (SAA) to solve the parametric optimization problem associated with the optimal importance sampling change of measure in simulating Lévy processes. Numerical experiments on variance gamma (VG), geometric Brownian motion (GBM), and normal inverse Gaussian (NIG) examples illustrate the computational advantages of the SAA-Newton algorithm over stochastic approximation (SA) based algorithms.

### 1 INTRODUCTION

Lévy processes are widely used to model financial underlying in derivatives pricing problems. Commonly used Lévy processes are the variance gamma (VG) process (Madan and Seneta 1990), the normal inverse Gaussian (NIG) process (Barndorff-Nielsen 1995), the generalized hyperbolic process (Eberlein and Keller 1995), the CGMY process (Carr et al. 2002), and the Meixner process (Schoutens and Teugels 1998). More complicated underlying models such as Lévy-driven stochastic volatility models have also been investigated in Barndorff-Nielsen and Shephard (2001) and Carr et al. (2003). For application of Greek estimation, see Fu (2007), Glasserman and Liu (2010).

In derivatives pricing, although analytical solutions are available for European-style options, most financial derivatives problems are solved by numerical methods, such as numerical solutions for PDEs (Cont and Voltchkova 2005), Fast-Fourier transition (FFT) (Carr and Madan 1999, Kwok et al. 2012), and Monte Carlo simulation (Glasserman 2003, Ribeiro and Webber 2006, Carr and Madan 2008). Although Monte Carlo simulation can solve high-dimensional problems, variance reduction techniques are often needed to improve the computational efficiency; see Glasserman (2003) for a survey. For Lévy processes models, variance reduction techniques used in derivatives pricing include importance sampling (Kawai 2008b), control variates (Dingçç and Hörmann 2012), stratified sampling (Kawai 2010), and their combination (Kawai 2008a).

In this paper, we focus on importance sampling. A commonly used approach is to formulate the importance sampling problem as a parametric optimization problem, and use stochastic approximation (SA) to solve this stochastic optimization problem. For more details, see Su and Fu (2002), Kushner and Yin (2003), Lapeyre and Lelong (2011). Kawai (2007) and Kawai (2008b) extended the approach to Lévy processes. SA is known to be sensitive to step sizes, which need to be chosen carefully. In this paper,

sample average approximation (SAA) (Shapiro 2003, Shapiro et al. 2009, Kim et al. 2014) is used to solve the resulting optimization problem. Jourdain and Lelong (2009) and Badouraly-Kassim et al. (2015) have used Newton's algorithm on SAA to find the optimal parameters in the Black-Scholes model and jump diffusion models, respectively. In our work, we consider more general Lévy process models.

The rest of the paper is organized as follows. In Section 2, we briefly review Lévy processes and the Monte Carlo method with importance sampling. Section 3 introduces the proposed SAA-Newton's method, and presents the resulting algorithms. Numerical examples are given in Section 4 for Asian options driven by normal inverse Gaussian (NIG) processes and geometric Brownian motion (GBM), and default probabilities of portfolios driven by variance gamma (VG) processes. Section 5 concludes.

## 2 LÉVY PROCESS MODELS AND MEASURE CHANGE

Lévy process models are commonly used to model the underlying assets. Let  $\mathbf{X} = \{X_t, t \geq 0\}$  be a  $d$ -dimensional stochastic process defined on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ , satisfying the following conditions: (i)  $X_0 = 0$  a.s. (ii)  $\mathbf{X}$  has independent and stationary increments. (iii)  $\mathbf{X}$  is stochastically continuous, i.e., for all  $a > 0$  and  $s \geq 0$ ,  $\lim_{t \rightarrow s} \mathbb{P}(\|\mathbf{X}_t - \mathbf{X}_s\| > a) = 0$ , where  $\|\cdot\|$  is the Euclidean norm. Then,  $\mathbf{X}$  is a Lévy process.

Let  $\Lambda$  denote the process parameters, e.g.,  $\Lambda = (\sigma, \nu, \theta)$  for the VG process and  $\Lambda = (\mu, \sigma)$  for Brownian motion, which determine the corresponding processes. Let  $F(X)$  be the payoff of the financial derivative, given by

$$F(X) = F(X_t; 0 \leq t \leq T),$$

where  $T$  is the maturity. The goal is to calculate the discounted expectation of  $F(X)$ .

To improve computing efficiency, variance reduction techniques are routinely used with Monte Carlo simulation. In this paper, we focus on importance sampling, which attempts to give more weight to "important" outcomes, thereby increasing sampling efficiency. In the Black-Scholes model, the drift is changed (Glasserman et al. 1999, Su and Fu 2002). For Lévy processes, the Esscher measure change is commonly used (Kawai 2008b).

Let  $\varphi_t(\lambda) = \log \mathbb{E}_{\mathbb{P}}[\exp(\langle \lambda, X_t \rangle)]$  denote the cumulant generating function of  $X_t$ , where  $\lambda \in C \subseteq \mathbf{R}^d$  and  $C$  is a nonempty convex compact set. Given another probability measure  $\mathbb{P}_\lambda$  that is absolutely continuous w.r.t.  $\mathbb{P}$ , the Radon-Nikodym derivative is given by

$$\left. \frac{d\mathbb{P}_\lambda}{d\mathbb{P}} \right|_{\mathcal{F}_t} = \frac{e^{\langle \lambda, X_t \rangle}}{\mathbb{E}_{\mathbb{P}}[e^{\langle \lambda, X_t \rangle}]} = e^{\langle \lambda, X_t \rangle - \varphi_t(\lambda)}, \quad (1)$$

where  $\mathcal{F}_t$  is the natural filtration of  $\{X_t, t \geq 0\}$ . Suppose that  $\mathbb{E}_{\mathbb{P}_\lambda}[F(X)] < \infty$ , then applying the Esscher measure change to  $\mathbb{E}_{\mathbb{P}}[F(X)]$ ,

$$V := \mathbb{E}_{\mathbb{P}}[F(X)] = \mathbb{E}_{\mathbb{P}_\lambda} \left[ \left. \frac{d\mathbb{P}}{d\mathbb{P}_\lambda} \right|_{\mathcal{F}_T} F(X) \right] = \mathbb{E}_{\mathbb{P}_\lambda} \left[ \left( \left. \frac{d\mathbb{P}_\lambda}{d\mathbb{P}} \right|_{\mathcal{F}_T} \right)^{-1} F(X) \right] = \mathbb{E}_{\mathbb{P}_\lambda} \left[ e^{-\langle \lambda, X_T \rangle + \varphi_T(\lambda)} F(X) \right].$$

The variance of  $F(X)$  under  $\mathbb{P}_\lambda$  is given by

$$\text{Var}(F(X), \lambda) := \mathbb{E}_{\mathbb{P}_\lambda} \left[ \left( \left. \frac{d\mathbb{P}}{d\mathbb{P}_\lambda} \right|_{\mathcal{F}_T} \right)^2 F(X)^2 \right] - V^2 = \mathbb{E}_{\mathbb{P}} \left[ e^{-\langle \lambda, X_T \rangle + \varphi_T(\lambda)} F(X)^2 \right] - V^2. \quad (2)$$

We call  $\lambda^*$  that minimizes the variance  $\text{Var}(F(X), \lambda)$ , the *optimal parameter*, i.e.,

$$\lambda^* \in \arg \min_{\lambda \in C} \text{Var}(F(X), \lambda). \quad (3)$$

Then we will provide an efficient algorithm to find  $\lambda^*$ . Note that solving problem (3) is equivalent to finding the parameters  $\lambda^*$  minimizing the second moment, i.e., finding

$$\lambda^* \in \arg \min_{\lambda \in C} \mathbb{E}_{\mathbb{P}} \left[ e^{-\langle \lambda, X_T \rangle + \varphi_T(\lambda)} F(X)^2 \right]. \quad (4)$$

### 3 THE SAA-NEWTON ALGORITHM AND IMPORTANCE SAMPLING

Under the SAA framework the optimization problem (4) is changed to a deterministic optimization problem. Denote

$$f(\lambda) := \mathbb{E}_{\mathbb{P}} [g(X, \lambda)],$$

where

$$g(X, \lambda) := e^{-\langle \lambda, X_T \rangle + \varphi_T(\lambda)} F(X)^2.$$

We generate independent and identically distributed (i.i.d.) paths of  $X$  denoted by  $X^1, X^2, \dots, X^n$ , and let

$$f_n(\lambda) = \frac{1}{n} \sum_{i=1}^n g(X^i, \lambda).$$

Then, we formulate a deterministic optimization problem

$$\min_{\lambda \in C} f_n(\lambda), \quad (5)$$

which can be solved by an iterative deterministic algorithm such as Newton's method. We call this approach the SAA-Newton algorithm. Generally, Newton's method is not globally convergent and may become computationally impractical in high dimensions due to the calculation of the Hessian matrix. However, in our setting, the SAA problem is convex (Jiang and Fu 2015), the problem dimension is low, and there is a natural starting point (parameter value 0), so Newton's method works well. For more details on the convergence properties of the algorithm, see Jiang and Fu (2015).

For the SAA-Newton algorithm,  $\nabla f_n(\lambda)$  and  $Hess[f_n(\lambda)]$  can be obtained by

$$\begin{aligned} \nabla f_n(\lambda) &= \frac{1}{n} \sum_{j=1}^n \left[ \left( \nabla \varphi_T(\lambda) - X_T^j \right) e^{-\langle \lambda, X_T^j \rangle + \varphi_T(\lambda)} F(X^j)^2 \right], \\ Hess[f_n(\lambda)] &= \frac{1}{n} \sum_{j=1}^n \left[ \left( Hess[\varphi_T(\lambda)] \right. \right. \\ &\quad \left. \left. + \left( \nabla \varphi_T(\lambda) - X_T^j \right) \left( \nabla \varphi_T(\lambda) - X_T^j \right)' e^{-\langle \lambda, X_T^j \rangle + \varphi_T(\lambda)} F(X^j)^2 \right) \right], \end{aligned}$$

which are used to find the optimal parameter in the Esscher measure change in Algorithm 1.

Similarly as in Jiang and Fu (2015), after obtaining the estimated  $\lambda_n^*$ , we can use the measure change equation to find the new process parameters set  $\Lambda^*$ . Algorithm 2 carries out importance sampling using the Esscher measure change.

## 4 NUMERICAL EXAMPLES

In this section, we provide numerical examples: pricing Asian call options driven by NIG process and GBM, and calculating the default probabilities of a portfolio of European call options driven by VG process.

### 4.1 Asian options

In this subsection, we compare the Newton-SAA algorithm with the SA approaches in Kawai (2008b) and Su and Fu (2002) in finding the optimal parameter of the Esscher measure change, and also estimate the resulting variance reduction in option pricing.

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**Algorithm 1** SAA-Newton algorithm for finding the optimal parameter

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**Input:** number of samples  $n$  in SAA; maturity  $T$ ; Lévy processes parameter  $\Lambda$  under the **original** probability measure; termination tolerance of Newton's method  $\rho$ .

**Initialization:** initial point  $\lambda_0$ ;  $k = 0$ .

- 1: generate and store Lévy process paths  $\{X_t^1, 0 \leq t \leq T\}, \{X_t^2, 0 \leq t \leq T\}, \dots, \{X_t^n, 0 \leq t \leq T\}$  under the **original** probability measure;
- 2: compute payoffs  $\{F(X^i), i = 1, 2, \dots, n\}$ ;
- 3: **repeat**
- 4:   compute  $\nabla f_n(\lambda_k)$  and  $Hess[f_n(\lambda_k)]$ ;
- 5:   solve  $Hess[f(\lambda_k)]\Delta\lambda_k = -\nabla f(\lambda_k)$ ;
- 6:   set  $\lambda_{k+1} = \lambda_k + \Delta\lambda_k$ ;
- 7:   set  $k = k + 1$ ;
- 8: **until** ( $\|\lambda_{k+1} - \lambda_k\| \leq \rho$ )

**Output:** estimated optimal parameter  $\lambda_n^* = \lambda_{k+1}$ .

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**Algorithm 2** Importance sampling by Esscher measure change

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**Input:** number of simulation samples  $N$  in pricing; maturity  $T$ ; Lévy processes parameter  $\Lambda$  under the **new** probability measure with the optimal parameter  $\lambda_n^*$ .

- 1: generate and store Lévy process paths  $\{X_t^1, 0 \leq t \leq T\}, \{X_t^2, 0 \leq t \leq T\}, \dots, \{X_t^N, 0 \leq t \leq T\}$  under the **new** probability measure;
- 2: compute  $\left\{V^i = F(X^i) e^{-\langle \lambda_n^*, X_t^i \rangle + \varphi_T(\lambda_n^*)}, i = 1, 2, \dots, N\right\}$ .

**Output:** estimated payoff  $\hat{V} = 1/N \sum_{i=1}^N V^i$ .

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#### 4.1.1 NIG process

Consider an Asian call option whose underlying asset  $S_t$  is driven by a NIG process, i.e.,

$$S_t = S_0 e^{at + X_t}, 0 \leq t \leq T,$$

where  $S_0$  is the initial price of the asset,  $T$  is the maturity, and  $a$  is a constant. If  $M$  is the number of observation points, and the averaging begins at 0, the payoff is given by

$$F(X) = \left( \frac{1}{M} \sum_{i=1}^M S_{t_i} - K \right)^+ = \left( \frac{1}{M} \sum_{i=1}^M S_0 e^{at_i + X_{t_i}} - K \right)^+,$$

where  $0 < t_1 < t_2 < \dots < t_M = T$  and  $\{X_{t_i}, i = 1, 2, \dots\}$  is a discrete NIG process. For more details, see Barndorff-Nielsen (1995), and Schoutens (2003).

The characteristic function for the NIG process is given by

$$\Phi_{X_t^{NIG}}(u) = \exp \left( -\delta t \left( \sqrt{\alpha^2 - (\beta + iu)^2} - \sqrt{\alpha^2 - \beta^2} \right) \right),$$

where  $t = 1$  is the characteristic function of a NIG random variable. The process parameters under the original probability measure are  $(\alpha, \beta, \delta)$ . According to Jiang and Fu (2015), the process parameters under the new probability measure are  $(\alpha, \beta + \lambda^*, \delta)$ , where  $\lambda^*$  is the optimal parameter.

We simulate the path of  $X_t$  using the independent increments and infinite divisibility of NIG random variables. Specifically, we generate an  $M$ -dimensional ( $d$  is replaced by  $M$  in the following) NIG random vector  $X' = (X'_1, X'_2, \dots, X'_M)$ , where  $X'_i$  is a NIG random variable with parameter  $\delta' = \delta/M$ , and let

$X_{t_i} = \sum_{j=1}^i X'_j$ . The payoff is then given by

$$F(X') = \left( \frac{1}{M} \sum_{i=1}^M S_0 e^{at_i + \sum_{j=1}^i X'_j} - K \right)^+. \tag{6}$$

Our objective is to find the optimal parameter  $\lambda^* = (\lambda_1^*, \lambda_2^*, \dots, \lambda_M^*)$ , where  $\lambda_i^*$  is the optimal parameter of  $X'_i$  in the measure change.

Based on Algorithms 1 and 2, computational results are presented. We consider process parameter values  $\Lambda = (\alpha, \beta, \delta) = (2, 0.2, 0.8)$  as in Lemaire and Pages (2009). For the other parameters, let  $a = 0$ , interest rate  $r = 0.02$ , initial price  $S_0 = 100$ , strike price  $K = 100$ , the number of observation points  $M = 5$ , maturity  $T = 1$ , and the initial value in Newton’s method  $\lambda_0 = [0 \dots 0]_{1 \times M}$ . First, we compare SA with the SAA-Newton algorithm. In SA, the step size is  $\varepsilon/m$ . In Fig.1, the left and right panels show the convergences of the norms of the gradient against the number of iterations for different step sizes  $\varepsilon$  in SA and different sample sizes  $n$  in SAA, respectively. Fig.1 illustrates SA’s well-known sensitivity to the choice of step size. On the other hand, for SAA larger sample size leads to better convergence rate, making it more straightforward for practitioners to apply. Although the SAA-Newton converges using far fewer iterations than SA, each iteration requires far more computation, so Table 1 compares total run time. The results indicate that the run time of the SAA-Newton algorithm is considerably less than SA for the same level of accuracy.

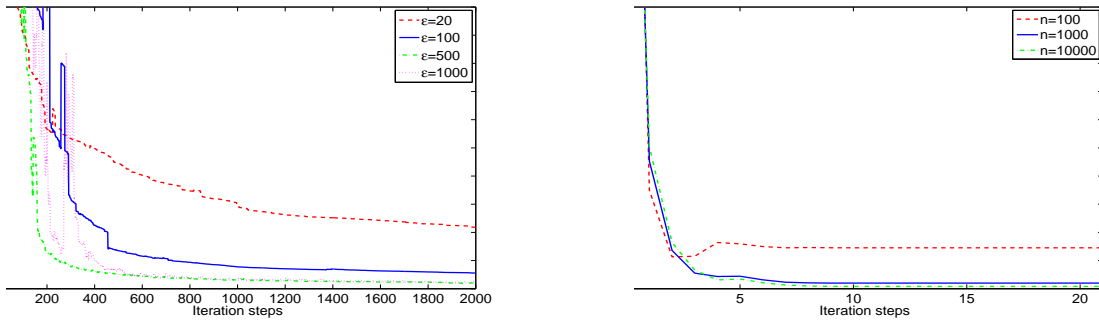


Figure 1: Norm of gradients for SA (left panel) and SAA-Newton (right panel)

Table 1: Run time and norm of gradient for SA and SAA-Newton

| $n$   | SAA-Newton |                  | $N$    | SA       |                  |
|-------|------------|------------------|--------|----------|------------------|
|       | run time   | norm of gradient |        | run time | norm of gradient |
| 250   | 0.054      | 0.0022           | 5000   | 0.27     | 0.016            |
| 500   | 0.11       | 0.0012           | 10000  | 0.54     | 0.0042           |
| 1000  | 0.21       | 0.00097          | 20000  | 1.1      | 0.0014           |
| 2500  | 0.52       | 0.00062          | 50000  | 2.7      | 0.0010           |
| 5000  | 1.0        | 0.00051          | 100000 | 5.4      | 0.00081          |
| 10000 | 2.1        | 0.00039          | 200000 | 10.8     | 0.00079          |
| 25000 | 5.2        | 0.00035          | 500000 | 27.0     | 0.00052          |

Next, we consider the importance sampling performance for the following cases: strike price  $K = 20, 100, 200$  and  $\delta = 0.4, 0.8$ . Figs. 2, 3, and 4 display the results in the form of box plots based on 100

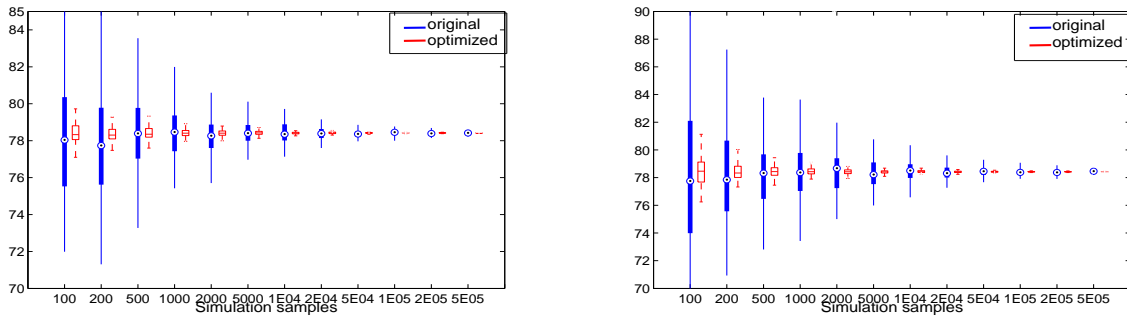


Figure 2: Convergence of prices for  $K = 20$ ;  $\delta = 0.4$  (left panel) and  $\delta = 0.8$  (right panel)

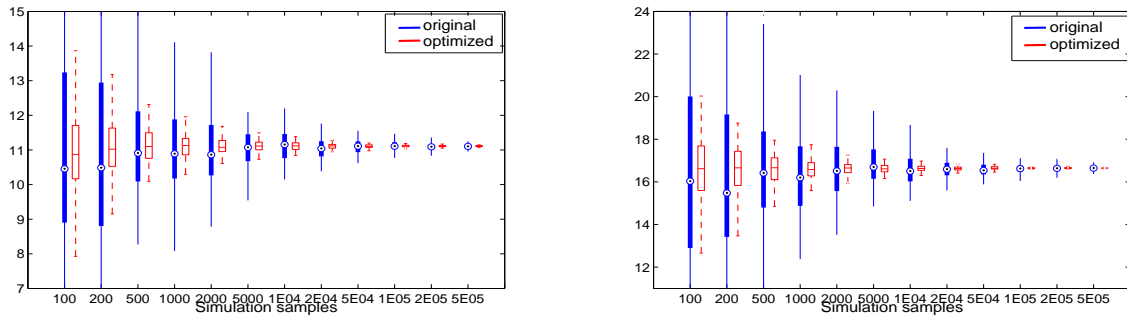


Figure 3: Convergence of prices for  $K = 100$ ;  $\delta = 0.4$  (left panel) and  $\delta = 0.8$  (right panel)

macro-replications. The results indicate that the optimal importance sampling reduces variance significantly, and Table 2 displays the variance reduction ratios defined as the variance of classical MC divided by the variance of MC with importance sampling. The results indicate significantly greater variance reduction for both out-of-the-money and in-the-money options than at-the-money options.

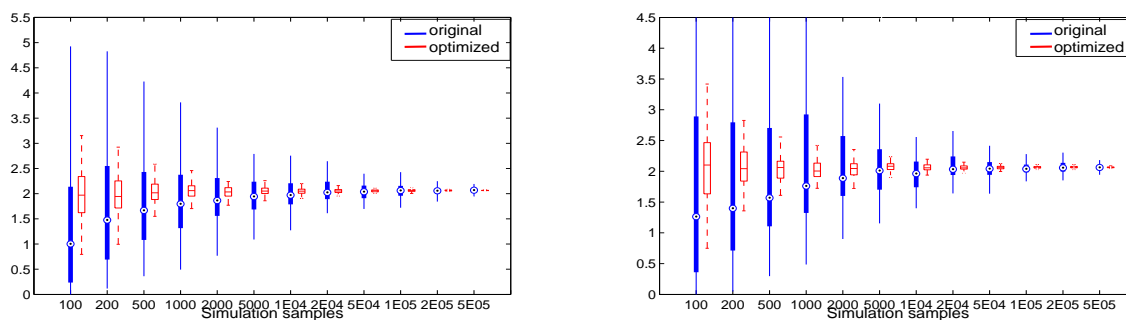
Table 2: Variance reduction ratios

| $N$    | $\alpha = 2, \beta = 0.2, \delta = 0.8$ |           |           | $\alpha = 2, \beta = 0.2, \delta = 0.4$ |           |           |
|--------|---|-----------|-----------|---|-----------|-----------|
|        | $K = 20$                                | $K = 100$ | $K = 200$ | $K = 20$                                | $K = 100$ | $K = 200$ |
| 500    | 60.9                                    | 14.1      | 68.2      | 52.8                                    | 11.9      | 72.5      |
| 1000   | 73.9                                    | 14.3      | 79.3      | 91.1                                    | 11.5      | 70.8      |
| 5000   | 57.7                                    | 14.3      | 64.8      | 67.5                                    | 14.1      | 70.1      |
| 10000  | 83.7                                    | 17.8      | 65.4      | 143.1                                   | 13.9      | 73.8      |
| 50000  | 78.1                                    | 14.2      | 55.2      | 56.2                                    | 14.3      | 101.2     |
| 100000 | 72.3                                    | 18.7      | 52.6      | 88.5                                    | 18.7      | 89.9      |

### 4.1.2 Geometric Brownian motion

Different from the SA approach in Kawai (2008b), Su and Fu (2002) proposed an SA approach based on estimated gradients via IPA. Consider an Asian call option whose underlying assets are driven by GBM, i.e.,

$$S_t = S_0 e^{(r - \frac{1}{2}\sigma^2)t + \sigma W_t},$$



**Figure 4: Convergence of prices for  $K = 200$ ;  $\delta = 0.4$  (left panel) and  $\delta = 0.8$  (right panel)**

where  $r$  is the interest rate,  $\sigma$  is volatility,  $S_0$  is the initial price of the asset, and  $\{W_t\}$  is standard Brownian motion under the original measure. Su and Fu (2002) change the drift in Brownian motion, which is a special case of the Esscher measure change. They assume a new Brownian motion

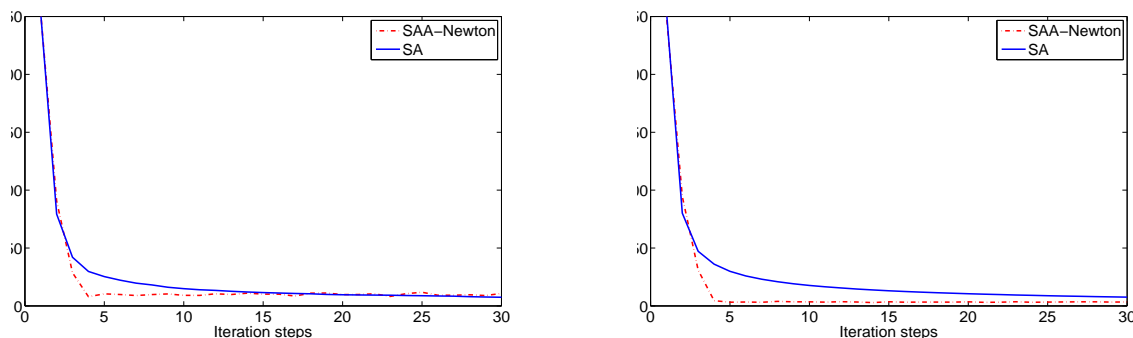
$$\tilde{W}_t = W_t - \lambda t,$$

with the Radon-Nikodym derivative given by

$$\frac{d\mathbb{P}_\lambda}{d\mathbb{P}} = e^{\lambda W_T - \frac{1}{2}\lambda^2 T}.$$

Note that, unlike the last example, we do not change the process parameter (the drift in Brownian motion) at each observation point, i.e., we use the same  $\lambda$ , which is a scalar, to obtain the new process parameter, at each observation point.

We consider the same setting in Su and Fu (2002):  $S_0 = 50$ ,  $K = 50$ ,  $\sigma^2 = 0.2$ ,  $r = 0.05$ ,  $T = 1$ , number of observation points  $M = 255$ , which is daily average and the averaging begins at 0. Let  $\tilde{n}$  denote the number of replications per iteration in SA. Fig.5 plots SAA-Newton and SA for  $\tilde{n} = n = 50$  and 500.



**Figure 5: Norm of gradients;  $\tilde{n} = n = 50$  (left panel) and  $\tilde{n} = n = 500$  (right panel)**

The results again indicate that SAA-Newton converges faster than SA. Table 3 shows the run times for SAA-Newton with iteration steps  $N = 15$  and SA with sample size  $\tilde{n} = 50$  per iteration. Although SAA-Newton requires more computation per iteration than SA, due to inverting a Hessian matrix, the overall run time of SAA-Newton is still less than SA for the same level of accuracy. Next, we consider the importance sampling performance. Similarly as in Su and Fu (2002), consider partial average Asian option pricing with the averaging beginning 60 days before the option maturity date. Let  $r = 0.05$ ,  $\sigma = 0.2$ ,  $S_0 = 100$ ,  $n = 50, 500$ , and  $K = 100, 120$ , with other parameter values remaining the same. Fig. 6 displays

Table 3: Run time, estimate drift and norm of gradient for SAA-Newton and SA

| $n$    | SAA-Newton |                |                | $N$    | SA       |                |                |
|--------|------------|----------------|----------------|--------|----------|----------------|----------------|
|        | run time   | estimate drift | $\ \nabla f\ $ |        | run time | estimate drift | $\ \nabla f\ $ |
| 50     | 0.08       | 0.488          | 11.5           | 10     | 0.04     | 0.417          | 16.2           |
| 500    | 0.13       | 0.494          | 2.8            | 100    | 0.48     | 0.475          | 4.0            |
| 5000   | 0.61       | 0.493          | 1.2            | 1000   | 4.58     | 0.489          | 1.2            |
| 50000  | 5.81       | 0.492          | 0.65           | 10000  | 45.5     | 0.492          | 0.82           |
| 500000 | 79.8       | 0.492          | 0.71           | 100000 | 458.4    | 0.493          | 0.742          |

the results in the form of box plots based on 100 macro-replications. The box plots depict the median, quartiles, and whiskers without the extreme points.

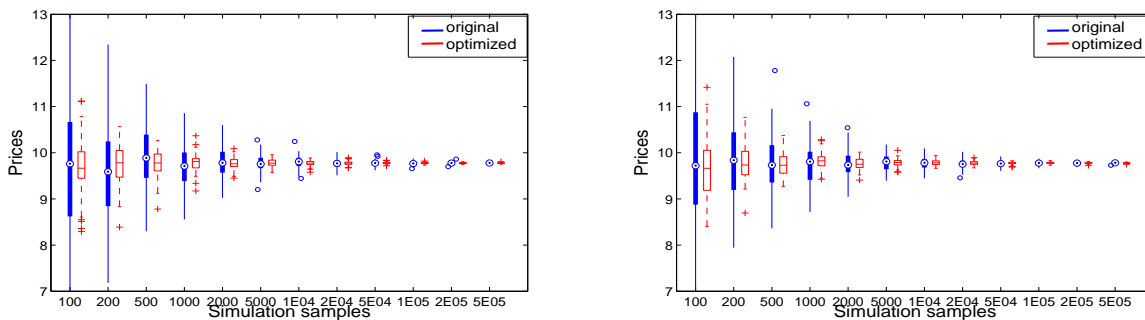


Figure 6: Convergence of prices for  $K = 100$ ;  $n = 50$  (left panel) and  $n = 500$  (right panel)

### 4.1.3 Comparison of two schemes for importance sampling

For the NIG example, we changed the process parameter at every observation point, so the importance sampling parameter is a vector, denoted here by IS-M. For the GBM example, we kept the drift the same, i.e., the optimal parameter is a scalar, denoted by IS-1.

We compare the variance reduction of these two schemes on Asian options driven by GBM, using SAA-Newton to find the optimal parameter. Let  $S_0 = 100$ ,  $K = 120$ ,  $T = 1$ ,  $M = 10$  averaged over entire period, iteration steps  $m = 20$ , with  $N = 10000$  for the pricing. Changing  $n$  w.r.t. three different values of  $(r, \sigma)$ , Table 4 is obtained. CMC stands for classical Monte Carlo, and VR-1 and VR-M stand for the variance reduction ratio for IS-1 and IS-M, respectively. IS-M has larger variance reduction ratios than the IS-1, but requires much longer run time than IS-1, which in this example outweighs the variance reduction. Also note that when  $n = 1000$ , the estimated optimal parameter for IS-1 is quite close to the true optimal parameter, so no further variance reduction is gained by increasing  $n$ . However, for IS-M,  $n = 1000$  is not enough to solve the high-dimensional problem, so higher  $n$  is required to better estimate the optimal parameter value.

### 4.2 Default probability of portfolio

In this subsection, we estimate the default probability of a simple portfolio  $L$  containing  $M$  European call options whose underlying assets are driven by VG processes, i.e.,

$$S_t^i = S_0^i e^{a_i t + X_t^i}, 0 \leq t \leq T, i = 1, 2, \dots, M, \tag{7}$$



Table 4: Comparison of IS-1 and IS-M

|                | CMC    |       | IS-1     |        |      | IS-M     |        |      |
|----------------|--------|-------|----------|--------|------|----------|--------|------|
|                | prices | $n$   | run time | prices | VR-1 | run time | prices | VR-M |
| $r = 0.15$     | 0.162  | 1000  | 0.01     | 0.165  | 9.6  | 0.18     | 0.160  | 16.0 |
| $\sigma = 0.1$ |        | 10000 | 0.06     | 0.159  | 10.5 | 1.79     | 0.161  | 35.3 |
| $r = 0.05$     | 0.712  | 1000  | 0.01     | 0.700  | 7.8  | 0.18     | 0.706  | 14.3 |
| $\sigma = 0.2$ |        | 10000 | 0.06     | 0.690  | 7.9  | 1.80     | 0.707  | 25.0 |
| $r = 0.05$     | 4.372  | 1000  | 0.01     | 4.313  | 5.6  | 0.18     | 4.362  | 12.4 |
| $\sigma = 0.4$ |        | 10000 | 0.06     | 4.353  | 5.8  | 1.82     | 4.368  | 16.7 |

where  $\{X_t^i\}$  is a VG process and  $\{S_0^1, S_0^2, \dots, S_0^M\}$  are the initial prices of the  $M$  underlying assets, and all the assets have the same maturity  $T$ , and  $\{a_i, i = 1, 2, \dots, M\}$  are constants. The portfolio is given by

$$L = \sum_{i=1}^M V_i,$$

with

$$V_i = e^{-rT} \mathbb{E}[(S_T^i - K_i)^+] = e^{-rT} \mathbb{E}[(S_0 e^{a_i + X_T^i} - K_i)^+],$$

where  $K_i$  is strike price of the  $i$ th option. The process parameters of  $X_T$  in the original probability are  $(\sigma, \nu, \theta)$ , according to Jiang and Fu (2015), the new process parameters for the Esscher measure change are

$$\left( \sigma / \sqrt{1 - \lambda \theta \nu - 1/2 \sigma^2 \nu \lambda^2}, \nu, (\theta + \lambda \sigma^2) / (1 - \lambda \theta \nu - 1/2 \sigma^2 \nu \lambda^2) \right).$$

Suppose we are interested in the probability that the portfolio is below some level  $L_b$ , i.e.,

$$\Pr\{L \leq L_b\} = \mathbb{E}[\mathbf{1}_{L \leq L_b}].$$

We know if  $L_b$  is small, this is the probability of a rare event. Let the process parameter  $(\sigma, \nu, \theta) = (0.1, 0.3, 0)$ , the number of options  $M = 5$ ,  $a_i = 0, i = 1, 2, \dots, M$ ,  $S_0 = [80 \ 90 \ 100 \ 110 \ 120]$  and the corresponding strike prices  $K = [70 \ 80 \ 90 \ 100 \ 110]$ . We consider sample sizes  $n = 500, 5000, 50000$  and default levels  $L_b = 10, 20$ . Fig.7 displays the convergence of the norm of the gradient via SAA-Newton. Similar to the previous examples, higher  $n$  leads to smaller norms of gradients.

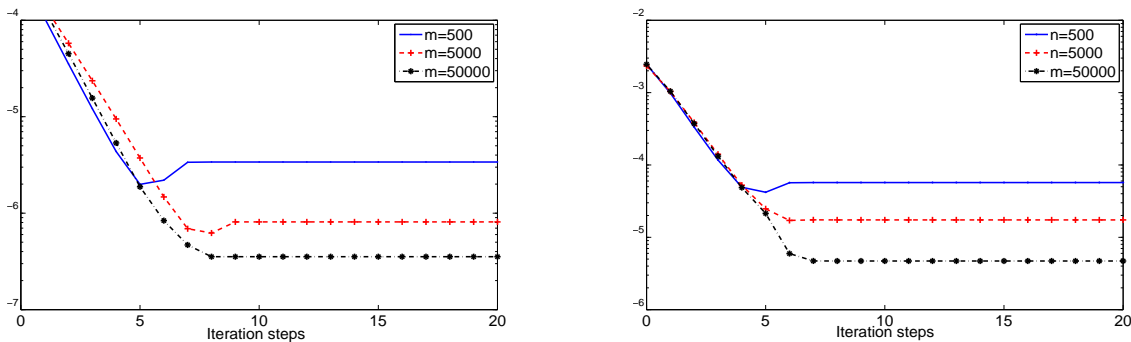
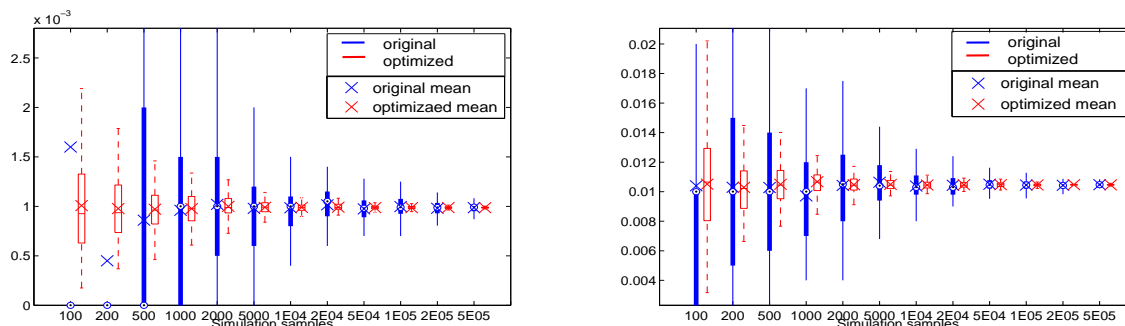


Figure 7: Norm of gradient;  $L_b = 10$  (left panel) and  $L_b = 20$  (right panel)

Next, we consider the importance sampling performance. Let  $n = 5000$ , and  $L_b = 10, 20$ , respectively, with other parameter values remaining unchanged. Fig. 8 displays the results in the form of box plots based

on 100 macro-replications. The cross marks in the box plots are the mean of the estimated probabilities of the 100 macro-replications. For  $L_b = 10$ , default is a rare event, and classical Monte Carlo returns an estimate of zero for small sample sizes ( $N = 100, 200$ ), whereas once importance sampling is applied, reasonable estimates can be obtained even for small  $N$ .



**Figure 8: Convergence of probabilities;  $L_b = 10$  (left panel) and  $L_b = 20$  (right panel)**

Tables 5 and 6 show the variance reduction, where  $P_c$  and  $P_{IS}$  denote the estimated probability using the classical Monte Carlo and the optimal importance sampling, respectively, with the 90% confidence half widths shown in parentheses. Clearly importance sampling works better for estimating rare event probabilities ( $L_b = 10$ ), especially for small sample size.

Table 5: Estimated probabilities (half widths) and variance reduction ratios for  $L_b = 10$

| $N$    | $P_c$             | $P_{IS}$          | VR    |
|--------|-------------------|-------------------|-------|
| 100    | 0.1600% (0.0667%) | 0.1008% (0.0082%) | 356.7 |
| 500    | 0.0860% (0.0219%) | 0.0971% (0.0037%) | 48.1  |
| 1000   | 0.0960% (0.0163%) | 0.0977% (0.0026%) | 42.5  |
| 5000   | 0.0980% (0.0074%) | 0.0990% (0.0012%) | 39.5  |
| 10000  | 0.0986% (0.0052%) | 0.0990% (0.0008%) | 39.5  |
| 50000  | 0.0975% (0.0023%) | 0.0988% (0.0004%) | 39.4  |
| 100000 | 0.0998% (0.0017%) | 0.0987% (0.0003%) | 40.2  |

Table 6: Estimated probabilities (half widths) and variance reduction ratios for  $L_b = 20$

| $N$    | $P_c$            | $P_{IS}$         | VR   |
|--------|------------------|------------------|------|
| 100    | 1.040% (0.1692%) | 1.051% (0.0489%) | 16.6 |
| 500    | 1.028% (0.0753%) | 1.050% (0.0217%) | 12.7 |
| 1000   | 0.972% (0.0518%) | 1.068% (0.0155%) | 11.5 |
| 5000   | 1.064% (0.0242%) | 1.050% (0.0069%) | 12.5 |
| 10000  | 1.038% (0.0169%) | 1.045% (0.0048%) | 12.2 |
| 50000  | 1.051% (0.0076%) | 1.047% (0.0022%) | 12.3 |
| 100000 | 1.045% (0.0054%) | 1.046% (0.0015%) | 12.3 |

## 5 CONCLUSION

In this paper, we first formulate the importance sampling problem as a parametric stochastic optimization problem, and then propose a new method, the SAA-Newton algorithm, to find the optimal importance sampling parameters based on the Esscher measure change for Lévy processes. Numerical experiments study the effectiveness of the method, and indicate that the Newton-SAA algorithm can find optimal parameters faster than SA-based algorithms.

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## AUTHOR BIOGRAPHIES

**GUANGXIN JIANG** is research assistant in the Department of Economics and Finance at City University of Hong Kong, Hong Kong. His research interests lie in simulation methodology, modeling, analysis, and optimization, with applications in financial engineering. His email address is [guajiang@cityu.edu.hk](mailto:guajiang@cityu.edu.hk).

**MICHAEL C. FU** is the Ralph J. Tyser Professor of Management Science in the Robert H. Smith School of Business, with a joint appointment in the Institute for Systems Research and affiliate appointment in the Department of Electrical and Computer Engineering, at the University of Maryland. His research interests include simulation modeling, analysis, and optimization, with applications to manufacturing, supply chain management, and financial engineering. He served as WSC 2011 Program Chair and is a Fellow of **INFORMS** and **IEEE**. His e-mail address is [mfu@rhsmith.umd.edu](mailto:mfu@rhsmith.umd.edu).

**CHENGLONG XU** is professor of Department of Mathematics at Tongji University, China. His research interests include simulation methodology, numerical analysis, spectral method, partial differential equations, with applications in financial engineering. His e-mail address is [clxu@tongji.edu.cn](mailto:clxu@tongji.edu.cn).